イロン イヨン イヨン イヨン 三日

1/38

Mathematical Background

Yuh-Jye Lee

Data Science & Machine Intelligence Lab Dept. of Applied Math @ NCTU

February 21, 2017



- 2 Multi-variable Calculus
- Probability and Statistics



Outline

1 Linear Algebra

2 Multi-variable Calculus

- Probability and Statistics
- Probability and Inference

< □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □

Norm

Definition

A norm is a function $\|\cdot\|:\mathbb{R}^n\to\mathbb{R}$ which must satisfy the following three conditions:

1
$$||x|| \ge 0$$
, and $||x|| = 0$ only if $x = 0$,

2
$$||x+y|| \le ||x|| + ||y||,$$

$$\|\alpha x\| = |\alpha| \|x\|.$$

Variants of Norm

- The most popular vector norms are defined below.
- The closed unit ball {x ∈ ℝⁿ : ||x|| ≤ 1} corresponding to each norm is illustrated to the right for the case n = 2.



Positive Definite Matrices

Definition

An $n \times n$ real symmetric matrix M is positive definite if $z^{\top}Mz > 0$ for all non-zero vectors $z \in \mathbb{R}^n$.

- Characteristics
 - All eigenvalues λ of M are positive.
 - There exists a unique lower triangular matrix L, with strictly positive diagonal elements, that allows the factorization of M into $M = LL^{\top}$. This factorization is called Cholesky decomposition.

Eigenvalues and Eigenvectors

Definition

Given a linear transformation A, a non-zero vector x is defined to be an eigenvector of the transformation if it satisfies the eigenvalue equation

$$Ax = \lambda x$$

for some scalar λ . In this situation, the scalar λ is called an eigenvalue of A corresponding to the eigenvector x.

• You can type *eigshow* in MATLAB to see the graphical demonstration of eigenvalues.



Diagonalization

A matrix $A_{n \times n}$ with *n* real eigenvalues $\lambda_1, \lambda_2, \ldots, \lambda_n$ and their associated eigenvectors q_1, q_2, \ldots, q_n can be diagonalized as follows:

$$A = Q \Lambda Q^{\top},$$

where

$$\Lambda = diag(\lambda_1, \lambda_2, \dots, \lambda_n), \ Q = [q_1|q_2| \dots q_n]$$

• The eigenvectors are the *principal components*. Extremely important in Machine Learning

Cholesky Factorization

- A matrix decomposition makes $A_{n \times n} = R_{n \times n}^{\top} R_{n \times n}$, where R is an upper-triangular matrix.
- The matrix A must be positive definite.

$$A = \begin{bmatrix} a_{11} & w^{t} \\ w & K \end{bmatrix}$$
(1)
$$= \begin{bmatrix} \alpha & 0 \\ \frac{w}{\alpha} & I \end{bmatrix} \begin{bmatrix} 1 & 0 \\ 0 & K - \frac{ww^{\top}}{a_{11}} \end{bmatrix} \begin{bmatrix} \alpha & \frac{w}{\alpha} \\ 0 & I \end{bmatrix}$$
(2)
$$= R_{1}^{\top} A_{1} R_{1}$$
(3)
$$= R_{1}^{\top} R_{2}^{\top} \cdots R_{m}^{\top} R_{m} \cdots R_{2} R_{1}$$
(4)
$$= R^{\top} R$$
(5)

QR Factorization

- A matrix decomposition makes $A_{m \times n} = Q_{m \times m} R_{m \times n}$, where $Q^{\top}Q = I_{m \times m}$ and R is an upper-triangular matrix.
- QR factorization can be computed by Gram-Schmidt process and Householder transformations.
 - Note: A matrix Q is called *orthogomal matrix* if $Q^{\top}Q = I$
- For a rectangular matrix:

$$A_{m \times n} = Q_{m \times m} R_{m \times n} = \begin{bmatrix} \hat{Q}_{m \times n} & Q_{m \times (m-n)}^{0} \end{bmatrix} \begin{bmatrix} \hat{R}_{n \times n} \\ 0 \end{bmatrix} = \hat{Q}_{m \times n} \hat{R}_{n \times n}$$

• $A = \hat{Q}\hat{R}$ is the reduced QR factorization

Singular Value Decomposition (SVD)

- A matrix decomposition makes $A_{m \times n} = U_{m \times m} \Sigma_{m \times n} V_{n \times n}^{\top}$, where $U^{\top} U = I_{m \times m}$ and $V^{\top} V = I_{n \times n}$.
- U and V are the eigenvectors of AA^{\top} and $A^{\top}A$ respectively.
- For a rectangular matrix:

$$A = U\Sigma V^{\top} = \begin{bmatrix} \hat{U}_{m \times n} & U^{0}_{m \times (m-n)} \end{bmatrix} \begin{bmatrix} \hat{\Sigma}_{n \times n} \\ 0 \end{bmatrix} V = \hat{U}_{m \times n} \hat{\Sigma}_{n \times n} V^{\top}_{n \times n}$$

- $A = \hat{U}\hat{\Sigma}V^{\top}$ is the reduced SVD.
- SVD is the Latent Semantic Indexing (LSI) in Text Mining when A is a *term by document* matrix

Least Squares Problem

• Given $A \in \mathbb{R}^{m \times n}$ and $y \in \mathbb{R}^m$, a linear system with m > n:

$$Aw = y, (6)$$

is called an overdetermined linear system.

 In general, an overdetermined linear system has no solution. An approximated solution can be obtained by solving the following minimization problem.

$$\min_{w \in \mathbb{R}^n} r^\top r = \min_{w \in \mathbb{R}^n} \|r\|_2^2 = \min_{w \in \mathbb{R}^n} \sum_{i=1}^m (y_i - A_i w)^2, \qquad (7)$$

where $r = y - Aw \in \mathbb{R}^m$ is the *residual*.

• The minimization problem (7) is the formulation of *least* squares problem.

Example: Data Fitting

Suppose we want to fit the data

$$(x_1, y_1), (x_2, y_2), \ldots, (x_m, y_m)$$

with a straight line $y = w_0 + w_1 x$.



・ロ ・ ・ 日 ・ ・ 目 ・ ・ 目 ・ う へ で
13/38

Example: Data Fitting

This problem can be expressed as the following overdetermined linear system:

$$y_1 = w_0 + w_1 x_1$$

$$y_2 = w_0 + w_1 x_2$$

$$\vdots$$

$$y_m = w_0 + w_1 x_m,$$

or

$$\begin{bmatrix} 1 & x_1 \\ 1 & x_2 \\ \vdots & \vdots \\ 1 & x_m \end{bmatrix} \begin{bmatrix} w_0 \\ w_1 \end{bmatrix} = \begin{bmatrix} y_1 \\ y_2 \\ \vdots \\ y_m \end{bmatrix}, \quad (8)$$

or

 A vector w minimizes the residual norm $||r||_2 = ||y - Aw||_2$, thereby solving the least squares problem if and only if $r \perp \operatorname{range}(A)$, that is,

$$A^{\top}r = 0$$

or equivalently,

$$A^{\top}Aw = A^{\top}y,$$

or equivalently,

$$Py = Aw,$$

where $P = A(A^{\top}A)^{-1}A^{\top}$ is a orthogonal projection and w is unique iff A is full rank ($w = (A^{\top}A)^{-1}A^{\top}y$).



Outline

1 Linear Algebra



- 3 Probability and Statistics
- Probability and Inference

・ロ ・ ・ 一部 ト ・ 注 ト ・ 注 ・ う へ で
16 / 38

Gradient

Definition

Let $f : \mathbb{R}^n \to \mathbb{R}$ be a differentiable function. The gradient of function f at a point $x \in \mathbb{R}^n$ is defined as

$$\nabla f(x) = \left[\frac{\partial f(x)}{\partial x_1}, \frac{\partial f(x)}{\partial x_2}, \dots, \frac{\partial f(x)}{\partial x_n}\right] \in \mathbb{R}^n$$

The gradient vector ∇f(x) gives the direction of fastest increase of f.

Example

$$f(x_1, x_2) = x_1^2 + x_2^2 - 2x_1 + 4x_2$$

$$\nabla f(x_1, x_2) = [2x_1 - 2 \quad 2x_2 + 4]$$

Hessian

Definition

If $f : \mathbb{R}^n \to \mathbb{R}$ is a twice differentiable function. The Hessian matrix of f at a point $x \in \mathbb{R}^n$ is defined as

$$\nabla^2 f(x) = \begin{bmatrix} \frac{\partial^2 f}{\partial x_1^2} & \frac{\partial^2 f}{\partial x_1 \partial x_2} & \cdots & \frac{\partial^2 f}{\partial x_1 \partial x_n} \\ \vdots & \vdots & \ddots & \vdots \\ \frac{\partial^2 f}{\partial x_n \partial x_1} & \frac{\partial^2 f}{\partial x_n \partial x_2} & \cdots & \frac{\partial^2 f}{\partial x_n^2} \end{bmatrix} \in \mathbb{R}^{n \times n}$$

- Hessian matrix describes the local curvature of a function
- Example

$$f(x_1, x_2) = x_1^2 + x_2^2 - 2x_1 + 4x_2$$

$$\nabla^2 f(x) = \begin{bmatrix} 2 & 0 \\ 0 & 2 \end{bmatrix}$$

18/38

Connection to Maximum and Minimum Values

First-Order Necessary Conditions

If x^* is a local minimizer and f is continuously differentiable in an open neighborhood of x^* , then $\nabla f(x^*) = 0$.

Second-Order Necessary Conditions

If x^* is a local minimizer and $\nabla^2 f$ exists and is continuous in an open neighborhood of x^* , then $\nabla f(x^*) = 0$ and $\nabla^2 f(x^*)$ is positive semidefinite.

Second-Order Sufficient Conditions

Suppose that $\nabla^2 f$ is continuous in an open neighborhood of x^* and that $\nabla f(x^*) = 0$, and $\nabla^2 f(x^*)$ is positive definite. Then x^* is a strict local minimizer of f.

イロト イポト イヨト イヨト

Revisit Least Squares Problem

• Given $A \in \mathbb{R}^{m \times n}$ and $y \in \mathbb{R}^m$, a linear system with m > n:

$$Aw = y, \tag{9}$$

is called an overdetermined linear system.

• Try to find an approximation solution with the "smallest residual"

$$\min_{w \in \mathbb{R}^n} \|r\|_2^2 = \min_{w \in \mathbb{R}^n} \sum_{i=1}^m (y_i - A_i w)^2 = \min_{w \in \mathbb{R}^n} f(w).$$
(10)

• Let $\nabla f(w) = \mathbf{0}$ we can have the *normal equation*

Outline

1 Linear Algebra

- 2 Multi-variable Calculus
- Probability and Statistics
- Probability and Inference

・ロ ・ ・ 一 ・ ・ 注 ・ ・ 注 ・ う へ で 21 / 38

Random Variable

Definition

A *random variable* is a real-valued function for which domain is a sample space

• Example

For a coin toss, the possible outcome is head or tail. The number of heads appearing in one fair coin toss can be described using the following random variable:

$$X = \left\{egin{array}{cc} 1, & ext{if head} \ 0, & ext{if tail} \end{array}
ight.$$

with probability function given by:

$$P(X = x) = \begin{cases} \frac{1}{2}, & \text{if } x = 1\\ \frac{1}{2}, & \text{if } x = 0\\ 0, & \text{sotherwise} \end{cases}$$

Probability Distribution

Definition

If X is discrete random variable, the function given by P(X = x) for each x within the range of X is called probability distribution of X.

• Example

Let the random variable X be denoted as the total number of heads. The probability distribution of heads obtained in the four tosses of a fair coin can be written as follows:

$$P(X = x) = \frac{\binom{4}{x}}{2^4}$$
, for $x = 0, 1, 2, 3, 4$.

◆□ → ◆□ → ◆目 → ◆目 → ○ へ ペ 23/38

Probability Density Distribution

Definition

A function with values f(x), defined over the set of all real numbers, is called a probability density function of the continuous random variable X if and only if

$$P(a \le X \le b) = \int_a^b f(x) dx,$$

for any real constants a and b with $a \leq b$

• Example

The p.d.f of normal distribution is defined as follows:

$$f(x) = \frac{1}{\sigma\sqrt{2\pi}}e^{-\frac{1}{2}(\frac{x-\mu}{\sigma})^2},$$

where μ is the mean and σ is the standard deviation.

24 / 38

< ∃ >

Conditional Probability

Definition

The conditional probability of an event A, given that an event B has occurred, is equal to

$$P(A|B) = rac{P(A \cap B)}{P(B)}$$

• Example

Suppose that a fair die is tossed once. Find the probability of a 1 (event A), given an odd number was obtained (event B).

$$P(A|B) = \frac{P(A \cap B)}{P(B)} = \frac{1/6}{1/2} = \frac{1}{3}$$

• Restrict the sample space on the event B

Theorem

Assume that $\{B_1, B_2, \ldots, B_k\}$ is a partition of S such that $P(B_i) > 0$, for $i = 1, 2, \ldots, k$. Then

$$P(A) = \sum_{i=1}^{k} P(A|B_i)P(B_i).$$



26 / 38

Note that {B₁, B₂,..., B_k} is a partition of S if
 S = B₁ ∪ B₂ ∪ ... ∪ B_k
 B_i ∩ B_j = Ø for i ≠ j

Bayes' Rule

Bayes' Rule

Assume that $\{B_1, B_2, \ldots, B_k\}$ is a partition of S such that $P(B_i) > 0$, for $i = 1, 2, \ldots, k$. Then

$$P(B_j|A) = \frac{P(A|B_j)P(B_j)}{\sum_{i=1}^k P(A|B_i)P(B_i)}.$$



Expected Value

Definition

If X is a discrete random variable and P(X = x) is the value of its probability distribution at x, the expected value of X is

$$\mu = E(X) = \sum_{x} x \cdot P(X = x).$$

Correspondingly, if X is a continuous random variable and f(x) is the value of its probability density at x, the expected value of X is

$$E(X)=\int_{-\infty}^{\infty}x\cdot f(x)dx.$$

• E(aX + bY) = aE(X) + bE(Y), linear operator

< □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □

Variance Measures of how far a set of numbers are spread out

Definition

If X is a discrete random variable and P(X = x) is the value of its probability distribution at x, the expected value of X is

$$Var(X) = E([X - E(X)]^2) = \sum_{x} (x - \mu)^2 \cdot P(X = x).$$

Correspondingly, if X is a continuous random variable and f(x) is the value of its probability density at x, the expected value of X is

$$Var(X) = \int_{-\infty}^{\infty} (x-\mu)^2 \cdot f(x) dx.$$

•
$$Var(X) = E(X^2) - (E(X))^2$$

< □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □

Bernoulli Distribution

A trial is performed whose outcome is either a "success" or a "failure". The random variable X is a 0/1 indicator variable and takes the value 1 for a success outcome and is 0 otherwise. p is the probability that the result of trail is a success. Then

$$P(X = 1) = p$$
 and $P(X = 0) = 1 - p$

which can equivalently be written as

$$P(X = i) = p^{i}(1-p)^{1-i}, i = 0, 1$$

Tossing a *fair* coin, the parameter p = 0.5. If X is Bernoulli,

$$\bullet E(X) = p,$$

2
$$Var(X) = p(1-p)$$

Who knows p?

Probability and Inference

- The outcome of tossing a coin is {*Heads*, *Tails*}
- We use a random variable $X \in \{0, 1\}$ to indicate the outcome
- Suppose that we have a random sample: $\mathbf{X} = \{x^t\}_{t=1}^N$
- How to *estimate* the parameter *p*?

Maximum Likelihood Estimation

Likelihood Function

The probability to *observe* the random sample $\mathbf{X} = \{x^t\}_{t=1}^N$ is

$$\prod_{t=1}^N p^{x^t}(1-p)^{1-x^t}$$

Why don't we choose the parameter p which will maximize the probability for observing the random sample $\mathbf{X} = \{x^t\}_{t=1}^N$?

Based on MLE, we will choose the parameter p

$$\rho = \frac{\sum_{t=1}^{N} x^t}{N}$$

Sample Mean, Variance, and Standard deviation

Sample Mean

The mean of a sample of *n* measured responses y_1, y_2, \ldots, y_n is given by

$$\bar{y} = \frac{1}{n} \sum_{i=1}^{n} y_i.$$

The corresponding population mean is denoted by μ .

Sample Variance

The variance of a sample of measurements y_1, y_2, \ldots, y_n is given by

$$s^2 = \frac{1}{n-1} \sum_{i=1}^n (y_i - \bar{y})^2.$$

The corresponding population variance is denoted by σ^2 .

Applying Baye's Rule to Classification

Credit Cards Scoring: Low-risk vs. High-risk

- According to the past transactions, some customers are low-risk in that they paid back their loan and the bank profited from them and other customers are high-risk in that they defaulted.
- We would like to learn the class "high-risk customer"
- We observe customer's *yearly income* and *savings*, which we represent by two *random variables* X_1 and X_2
- The *credibility of a customer* is denoted by a *Bernoulli* random variable *C* where *C* = 1 indicates a high-risk customer and *C* = 0 indicated a low-risk customer

Applying Baye's Rule to Classification

How to make the decision when a new application arrives?

- When a new application arrives with $X_1 = x_1$ and $X_2 = x_2$
- If we know the probability of *C* conditioned on the observation $X = [x_1, x_2]$ our decision will be

•
$$C = 1$$
 if $P(C = 1 | [x_1, x_2]) > 0.5$

- C = 0 otherwise
- The probability of error we made based on this rule is

$$1 - \max\{P(C = 1 | [x_1, x_2]), P(C = 0 | [x_1, x_2])\} < 0.5$$

• Please note
$$P(C = 1|[x_1, x_2]) + P(C = 0|[x_1, x_2]) = 1$$

The Posterior Probability: $P(C|\mathbf{x}) = \frac{P(C)P(\mathbf{x}|C)}{P(\mathbf{x})}$

- P(C = 1) is called the *prior probability* that C = 1
- In our example, it corresponds to a probability that a customer is high-risk, *regardless* of the x value.
- It is called the *prior probability* because it is the knowledge we have *before* looking at the observation **x**
- $P(\mathbf{x}|C)$ is called the *class likelihood* and is the *conditional probability* that an *event belonging to the class C* has the associated observation value \mathbf{x}
- *P*(**x**), the *evidence* is the probability that an observation **x** to be seen, regardless of whether it is a positive or negative example

All above information can be extracted from the past transactions (historical data)

The Posterior Probability: $P(C|\mathbf{x}) = \frac{P(C)P(\mathbf{x}|C)}{P(\mathbf{x})}$

- Because of normalization by the evidence, the posteriors sum up to 1
- In our example, P(X₁, X₂) is called the *joined probability* of two random variables X₁ and X₂
- Under the assumption, these two random variables X₁ and X₂ are *probability independent*, we have P(X₁, X₂) = P(X₁)P(X₂)
- It is one of key assumptions of Naive Bayes' Classifier
- Although it is *over simplified* the problem it is very easy to use for real applications

Extend to Multi-class classification

- We have *K* mutually and exhaustive classes; *C_i*, *i* = 1, 2, ..., *K*
- For example, in *optical digit recognition*, the input is a *bitmap image* and there are 10 classes
- We can think of that these *K* classes define a *partition* of the *input space*
- Please refer to the slides of the *Partition Theorem* and *Baye's Rule*
- The Bayes' classifier choose the class with the highest posterior probability; that is we choose *C_i* if

$$P(C_i|\mathbf{x}) = \max_k P(C_k|\mathbf{x})$$

• Question: Is it very important to have $P(\mathbf{x})$, the evidence?

38 / 38